

# Perceived value of video games, but not hours played, predicts mental well-being in adult Nintendo players

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Studies on video games and well-being often rely on self-report measures or data from a single game. Here, we study how 703 US adults' time spent playing for over 140,000 hours across 150 Nintendo Switch games relates to their life satisfaction, affect, depressive symptoms, and general mental well-being. We replicate previous findings that playtime over the past two weeks does not predict well-being, and extend these findings to a wider range of timescales (one hour to one year). Results suggest that relationships, if present, dissipate within two hours of gameplay. Our non-causal findings suggest substantial confounding would be needed to shift a meaningful true effect to the observed null. Although playtime was not related to well-being, players' assessments of the value of game time—so called gaming life fit—was. Results emphasise the importance of defining the gaming population of interest, collecting data from more than one game, and focusing on how players integrate gaming into their lives rather than the amount of time spent.

*Keywords:* video games, digital trace data, well-being

1 Video games played on smartphones, computers, or home  
2 consoles are now a widely pursued form of leisure that in-  
3 volves social interaction, creativity, problem-solving, and  
4 growth (Bourgonjon et al., 2016). Major firms like Nintendo  
5 have sold hundreds of millions of games consoles in recent  
6 years (Nintendo Japan, 2024) and online platforms such as  
7 Steam regularly attract upwards of 30 million players online  
8 at any given time (SteamCharts, 2024). This staggering in-  
9 vestment of human attention and behaviour has inspired both  
10 national (American Psychiatric Association, 2013) and inter-

national (World Health Organization, 2018) health bodies to  
focus on play as potential contributor to psychopathology.

The extent to which games might be understood as be-  
haviorally addictive remains hotly debated (Aarseth et al.,  
2017; Billieux et al., 2017; Van Rooij et al., 2018) and the  
broader scientific conversation has increasingly focused on  
how not just quantity, but also quality of play relates to player  
health. Although there is increasing recognition that not all  
screentime—or in the case of games, playtime—is created  
equal (Orben, 2022), it remains a major research focus. Re-  
search has identified a range of factors that co-determine how  
time spent with games relates to health: for example, as-  
pects of a game's design such as its social affordances (Cren-  
shaw & Nardi, 2016), the context of when and where one  
plays (Drummond & Sauer, 2020), and players' motivation  
(Brühlmann et al., 2020).

In matters of health policy, overall time spent with  
games—regardless of what or why one plays—remains cen-  
tral to how games are thought to influence player out-  
comes. Parental control tools on platforms like Xbox and  
PlayStation foreground time limits as a primary means of  
enforcing healthy gaming behavior (Robertson, 2021); for  
adults, a growing array of self-control tools, apps, and dash-  
boards offer the ability to “limit or cut yourself off from  
apps and games” (<https://focusme.com>), savings users “1.23  
hours” (<https://www.opal.so>) or “up to 2.5 hours a day”  
(<https://freedom.to/>). News media suggest time-based limits  
(e.g., Saveva, 2023), often referencing the since-abandoned

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2x2 rule from the American Academy of Pediatrics: no screen time for children under two, and no more than two hours per day for children older than two (Blum-Ross & Livingstone, 2018). Likewise, the American Psychiatric Association’s description of Internet Gaming Disorder characterises pathological engagement with games, in part, as involving ‘8-10 hours or more per day [and or] least 30 hours per week’ (American Psychiatric Association, 2013, p. 796). On a larger scale, time-based limits such as Korea’s 10-year Youth Protection Revision Act prohibited young people from playing games between 00:00 and 06:00 (Sang et al., 2017). More recently, China put in place a weekly three-hour limit for under-18s (Feiner & Kharpal, 2021). The effectiveness of such regulatory steps has been challenged (Choi et al., 2018; Zendle et al., 2023).

A better understanding of how time spent with games relates to players’ health, for good or ill, is needed. Given that play takes many forms and happens across many different games, researchers greatly benefit from access to digital trace data—histories of user actions generated when interacting with technologies such as a game or online platform. Digital trace data can provide much greater detail about what, when, and how much people play than is possible in self-report data, which consistently shows substantial discrepancies compared to digital trace data collected by online platforms (Ernala et al., 2020) and independent researchers (Parry et al., 2021) alike. Previous studies on games found that an additional hour of *Animal Crossing: New Horizons* trace data predicted just a 30-minute increase in self-reported play—a nearly 50% discrepancy (Johannes et al., 2021)—and that *Everquest 2* players’ estimates correlated only  $r = .37$  with logged estimates, with underestimates slightly more common than overestimates (Kahn et al., 2014).

Only a handful of studies have applied digital trace data to the study of games and well-being (Brühlmann et al., 2020; Johannes et al., 2021; Larrieu et al., 2023; see Vuorre et al., 2022), in part because this data can be very difficult to acquire: researchers must build or rely on unstable technical systems to log data themselves, or negotiate individual agreements with games companies who have historically been reluctant to share data (Ballou, 2023; Seif El-Nasr et al., 2013). Where digital trace data has been collected, however, results have been informative. Brühlmann et al. (2020) used playtime and in-game behavior to identify subgroups of *League of Legends* players who had more negative in-game experiences. Johannes et al. (2021) look at playtime in *Animal Crossing: New Horizons* and *PvZ: Battle for Neighborville* and found a positive but likely negligible correlation. A follow-up study expanded this to seven games, finding that changes in playtime over the course of 6 weeks were unlikely to affect subsequent well-being (Vuorre et al., 2022). Larrieu et al. (2023) follow high-intensity *Rainbow 6: Siege* players and find no link between playtime and quality of life across any identified

player types.

Though informative, this early work has a vital scope limitation: digital trace data was only available for a single game for each player-participant. It was not possible to know what other games a participant was, or was not, electing to play. Players regularly switch between games over the course of a day or week based on mood, available time, and social context (Ballou et al., 2024; O’Neill et al., 2016); data collected for one particular game may thus tell us little about *overall* playtime or its relation to well-being.

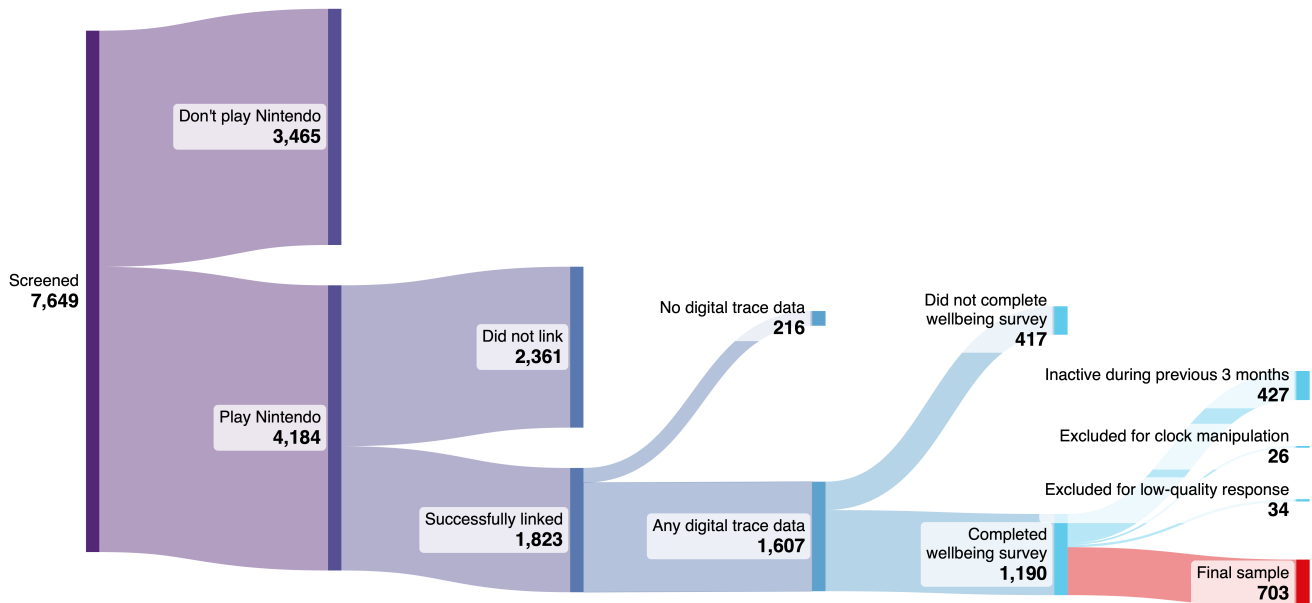
An important frontier for the field, therefore, is to collect holistic digital trace data that reflects behavior in not just one game, but all games played (which may include various games on one platform, such as Nintendo Switch or Steam, or all games played across several platforms a player uses). At present, our understanding of even basic phenomena such as the true volume of play in different demographics rely on the same inaccurate self-report data, itself often provided by market research firms using opaque methodologies. Capturing play at the platform level represents one step towards this goal of studying a player’s complete gaming diet. To our knowledge, the only study to platform-level digital trace data to investigate player health is (Ballou et al., 2024), which found no meaningful relationships between Xbox playtime and well-being over three months.

## Present Study

In the present research we report on a study conducted in collaboration with Nintendo of America, in which we independently recruited a large sample of adult video game players, surveyed them about their motivation and well-being, and joined these responses to digital trace data on Nintendo Switch video game play. Our central aim was to test the extent to which the amount of time participants spent playing related to their psychological well-being. Our analysis plan was preregistered at <https://osf.io/sjqyt>.

More specifically, our first hypothesis was to test whether the null relations reported in earlier work (Ballou et al., 2024; Johannes et al., 2021; Vuorre et al., 2022) would replicate in an independent sample of Nintendo Switch play. In our first hypothesis, we predicted that there will be no practically significant association between video game playtime in the previous 2 weeks and life satisfaction (H1a), affect (H1b), depressive symptoms (H1c), or general mental well-being (H1d).

We were also interested in understanding how relationships between play and well-being might depend on the choice of what timescale of play researchers consider relevant to investigate. . To this end, we examined relationships between well-being and a wide range of time windows of play preceding the well-being question. While preregistered, the large number of models implicated in our analysis plan is prohibitively large to be aptly framed as a narrow hypothe-

**Figure 1***Participant flow, from recruitment to final sample*

144 sis test. We therefore operationalise this as an exploratory<sup>168</sup>  
 145 research question: How does the relationship between play<sup>169</sup>  
 146 time and well-being differ across different gaming observa<sup>170</sup>  
 147 tion windows ranging from one hour to one year?<sup>171</sup>

148 Lastly, we investigated what factors might moderate the<sup>172</sup>  
 149 relationship between playtime and well-being. To test this ex<sup>173</sup>  
 150 ploratory question, we assessed potential moderation by par<sup>174</sup>  
 151 ticipant gender and age, as well as a subjective sense of how<sup>175</sup>  
 152 players thought gaming had positive or negative relationships<sup>176</sup>  
 153 to various life domains such as work, relationships, school<sup>177</sup>  
 154 performance, and social health.<sup>178</sup>

## 155 Method

### 156 Design and Recruitment

157 Our participant flow is shown in Figure 1. We recruited<sup>183</sup>  
 158 participants from Prolific, a participant recruitment platform,<sup>184</sup>  
 159 who were: (1) 18 years old, (2) proficient in English, (3)<sup>185</sup>  
 160 residents of the US, and (4) active video game players (self-<sup>186</sup>  
 161 defined, based on a >0 response to Prolific's built-in screening<sup>187</sup>  
 162 item "How many hours per week do you play video games on<sup>188</sup>  
 163 average?"). We first distributed a screening questionnaire to<sup>189</sup>  
 164 7649 participants asking which video game platforms they<sup>190</sup>  
 165 were active on; of these, 4184 indicated that they played<sup>191</sup>  
 166 games on Nintendo Switch.<sup>192</sup>

167 We invited these participants to a second survey wherein<sup>193</sup>

they retrieved an account identifier from the Nintendo web  
 interface using the events QR code, available at [https://  
 accounts.nintendo.com/qrcode](https://accounts.nintendo.com/qrcode), and they shared these unique  
 identifiers with us. This identifier is separate from their user-  
 name, and cannot be used by anyone besides Nintendo—  
 including our research team—to link the player to personally  
 identifiable information. A total of 1823 participants com-  
 pleted the linking process. We sent the account identifiers  
 to Nintendo of America, who in turn sent us each player's  
 pseudonymised play history from May 1, 2022 to present.  
 Data collection began with a pilot study of 100 participants  
 on Nov 15, 2023, which was combined with primary data col-  
 lection from Feb 12, 2024 to May 6, 2024.

181 Of the participants who completed the linking process,  
 182 1607 had eligible Nintendo data (i.e., any play sessions of  
 a 1st party game—a game published by Nintendo or one of  
 its subsidiaries, as opposed to a third-party publisher—from  
 May 1, 2022 to present). These 1607 were subsequently sent  
 a well-being survey in Qualtrics detailed below. 1191 par-  
 ticipants completed the well-being survey; our preregistered  
 stopping rule went into effect when 5 or fewer participants  
 per day completed the survey for three consecutive days.

As preregistered, we excluded 427 individuals who had  
 no playtime during the previous three months, indicating that  
 they are not active Nintendo players, and 26 people who  
 logged more than 24 hours of playtime on any single day

194 or who had a session that took place in the future, indicat-245  
 195 ing a technical problem or manipulation of the system clock246  
 196 for in-game benefits. We further excluded 34 participants247  
 197 who displayed signs of careless responding (e.g., so-called248  
 198 straightlining or seemingly random responses). In total, we249  
 199 excluded 487 participants (some participants were excluded250  
 200 on multiple grounds), leading to a final sample of 703. 251

201 Participants were paid £0.15 for the 1-minute screen-252  
 202 ing questionnaire, £3 for linking their data plus a £2253  
 203 bonus payment if data was successfully retrieved, and £4254  
 204 for a 22-minute well-being survey. This study received255  
 205 approval from the University of Oxford Social Sciences256  
 206 and Humanities Interdivisional Research Ethics Committee257  
 207 (OII\_C1A\_23\_107). 258

## 208 Participants and Exclusions 259

209 Participant demographics are shown in Table 1. 260

## 210 Measures 261

211 Participants completed a self-report survey that took on 262  
 212 average 22 minutes to complete. The survey included back-263  
 213 ground factors such as demographics and life circumstances,264  
 214 a series of cognitive tasks, as well as self-report measures of265  
 215 time use, and motivations for video game play. We detail266  
 216 those measures we used in this study below but all study data267  
 217 are available at OSF (see supplementary materials). 268

218 **Video game playtime** was measured by collecting a record269  
 219 of each player's individual game sessions for all 1st party270  
 220 video games played on a Nintendo Switch. These data were271  
 221 provided by Nintendo of America. Playtime was calculated272  
 222 by summing the duration of all (partial) sessions that oc-273  
 223 cur during a given time period relative to the participant's274  
 224 survey response, based on the logged session start and end275  
 225 times. For ease of interpretation, game play time in all obser-276  
 226 vation periods longer 24 hours is reported as mean minutes277  
 227 of play per day. It is important to note that these data in-278  
 228 cluded only telemetry for games published by Nintendo and279  
 229 its close business partners (e.g. The Pokémon Company), but280  
 230 not games made by third party publishers (e.g. Electronic281  
 231 Arts). Nintendo-published games accounted for 63% of all282  
 232 playtime across our sample; the 37% of play data from 3rd-283  
 233 party games is therefore treated as missing. We return to this284  
 234 limitation in the discussion. 285

235 **General mental well-being** was measured with the286  
 236 WEMWBS (Tennant et al., 2007). Players rated 14 state-287  
 237 ments about how they felt during the past 2 weeks such as288  
 238 "I've been dealing with problems well" and "I've been feel-289  
 239 ing good about myself" on a 5-point scale from 1 ("none of290  
 240 the time") to 5 ("all of the time"). Scores were calculated by291  
 241 taking the mean of all items. 292

242 **Depressive symptoms** was measured with the PROMIS293  
 243 Short Form 8a (Pilkonis et al., 2011). Players rated 8 state-294  
 244 ments about how they felt in the past 7 days such as "I felt295

hopeless" and "I felt I had nothing to look forward to" on a  
 5-point scale from 1 ("never") to 5 ("always"). Scores were  
 calculated by taking the mean of all items.

**Life satisfaction** was measured with the one-item Cantril  
 self-anchoring scale (Cantril, 1965). Participants were  
 prompted with 'Please imagine a ladder with steps numbered  
 from 0 at the bottom to 10 at the top. The top of the ladder  
 represents the best possible life for you, and the bottom of the  
 ladder represents the worst possible life for you. On which  
 step of the ladder would you say you personally feel you stood  
 over the past two weeks?'. Participants responded on a scale  
 from 0 to 10, which was rescaled to 1-5 to match the other  
 well-being measures.

**Affect** was measured with a single item: "How are you  
 feeling right now?" (Killingsworth & Gilbert, 2010). Par-  
 ticipants responded using a 100-point visual analogue scale  
 (VAS) with endpoints "Very bad" and "Very good", which  
 we rescaled to 1-5 to match the other well-being measures.

**Gaming life fit** was measured with a draft measure asking  
 players to rate the contribution of gaming to 5 life domains  
 (work/school, social participation, cognitive health, emotion  
 regulation, and daily routines) on a 7-pt bipolar scale from  
 "greatly interfered" to "greatly supported". We took the aver-  
 age of these to generate a formative indicator of the degree to  
 which players perceive their gaming to be beneficial or harm-  
 ful to other aspects of their lives. This measure has not been  
 used or validated before, and we return to this in the discus-  
 sion.

## 235 Analytic Approach 286

236 To test H1, we fit multiple regression models with playtime  
 237 over the previous 2 weeks as the primary predictor, all de-  
 238 mographic variables as covariates (age, gender, highest level  
 239 of education, and employment status), and the corresponding  
 240 mental health variable as the outcome. For example, for H1a  
 241 (life satisfaction predicted by the previous 2 weeks of play),  
 242 the model in R is:

```
243 {{lm((life satisfaction) ~ (playtime in the previous 2  

  244 weeks) + age + gender + education + employment)}}
```

245 We apply a similar analysis approach to our exploration of  
 246 H2 concerning other timescales; we primarily apply multi-  
 247 ple regression with well-being predicted by playtime aggre-  
 248 gated over various time periods and the same covariates, but  
 249 also explore potential non-linear alternatives and moderation  
 250 analyses (detailed below).

251 We interpret the playtime coefficient estimates from these  
 252 models in reference to pre-specified smallest effect sizes of  
 253 interest (see below): if the 99% confidence interval is fully  
 254 within the upper and lower equivalence bounds, this provides  
 255 evidence to reject a practically meaningful association.

256 We conducted all statistical analyses with R version 4.3.2  
 257 (R Core Team, 2023). We use an alpha of .01.

**Table 1***Participant Demographics*

Descriptor	Variable	Value
Age	Mean (SD)	31.5 (7.7)
	Median (IQR)	31.0 [26.0;36.0]
	Min / Max	18.0 / 68.0
Gender	Man	376 (53.5%)
	Woman	267 (38.0%)
	Non-binary or other gender identity	60 (8.5%)
Employment Status	Working full-time	338 (48.1%)
	Working part-time	120 (17.1%)
	Other employment status	97 (13.8%)
	Not currently employed	91 (12.9%)
	Student	57 (8.1%)
Ethnicity	White	456 (64.9%)
	Asian	82 (11.7%)
	Mixed	75 (10.7%)
	Black	54 (7.7%)
	Other ethnicity	36 (5.1%)

**Smallest Effect Size of Interest**

We specify the smallest effect size of interest (SESOI) as a 1-hour change in (daily) playtime associated with a .06 point change in mental health on a 1–5 scale, in line with Balou et al. (2024), who justified that value based on previous research on minimally important differences (approximately .3–.4 scale points on a 1–5 scale for PROMIS and WEMWBS measures) and daily leisure time available to US adults (approximately 5 hours; Sturm and Cohen (2019)). Any association smaller than .06 indicates that the average person does not have enough time in the day to modulate their play to an extent that it would register a perceptible difference in their well-being.

Note: this method of specifying an SESOI is predicated on a causal interpretation—it implicitly imagines a world where one can intervene on playtime (our predictor) and have an effect of a certain size on mental health (our outcomes). It is very unlikely that our cross-sectional analyses can provide unbiased causal estimates. Instead, our goal is to use associations to place boundaries on the size of a possible effect.

In other words, if there is no meaningful correlation between playtime and mental health, there is even less likely to be a meaningful causal effect between playtime and mental health. We support this reasoning with simulations presented in the discussion.

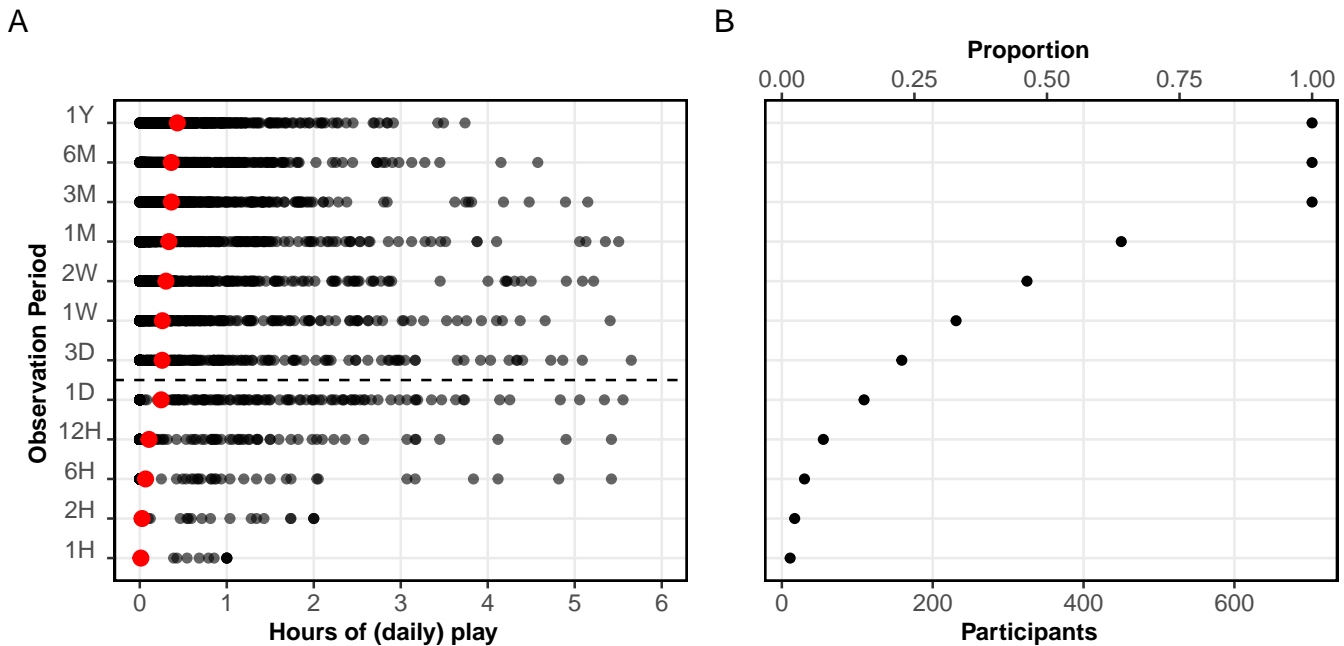
**Results****Descriptive results**

Given the lack of holistic or platform-level data available in the literature, our first goal was to simply describe the volume of play. This is visualised in Figure 2, which show that despite a total play volume of more than 140k hours, our sample was largely minimally engaged with 1st party Nintendo games. During the two weeks prior to survey completion, just over half of the sample had no sessions logged. The top 10% of players were moderately engaged, playing an average of 60 minutes per day. Sessions of a game lasted on average 41.9 minutes [10th percentile: 9.1; 90th percentile: 147.5].

The results of this study are therefore reflective of a largely casual population of players—at least with respect to Nin-

**Figure 2**

Description of playtime in the sample. Panel A shows the distribution of playtime across players for each observation period, with the mean shown in red; observations periods below the dashed line are shown as total hours, whereas periods above the dashed line reflect hours of player per day. Panel B shows the proportion of players who logged at least one play session in that period.



tendo titles. We argue that this population is nonetheless an important one: if video games were to meaningfully affect well-being, we may expect a larger impact for people who rarely play but happen to play for one hour, than for a highly engaged population of people who tend to play three hours per day, but happen to play four hours. We return to this limitation on generalizability in the discussion.

### H1: Previous 2 weeks of playtime and mental health

We began by analyzing H1, which concerned the relationship between mental health and the previous 2 weeks of playtime. This time period is common in the literature, and served as a way to conceptually replicate a previous study focused on one game (Johannes et al., 2021) using platform-level data.

Results are visualized in Figure 2. Multiple regression models found no evidence that people who played 1 additional hour per day in the previous 2 weeks differed from their peers with regard to life satisfaction ( $B = -0.02$  99% CI [-0.12, 0.05]), affect ( $B = 0.08$  99% CI [-0.03, 0.19]), depressive symptoms ( $B = -0.06$  99% CI [-0.19, 0.07]), or general mental well-being ( $B = 0.08$  99% CI [-0.02, 0.18]).

However, due to lower than expected response rates and to total volume of playtime, there is too much uncertainty around our estimates to confidently reject the presence of a meaning-

ful relationship using our original SESOI of .06; following our inference criteria, the results of our original hypothesis tests are all inconclusive. We therefore interpret our results as indicating an *absence of evidence* for a relationship between playtime and well-being, but do not conclude *evidence of absence*.

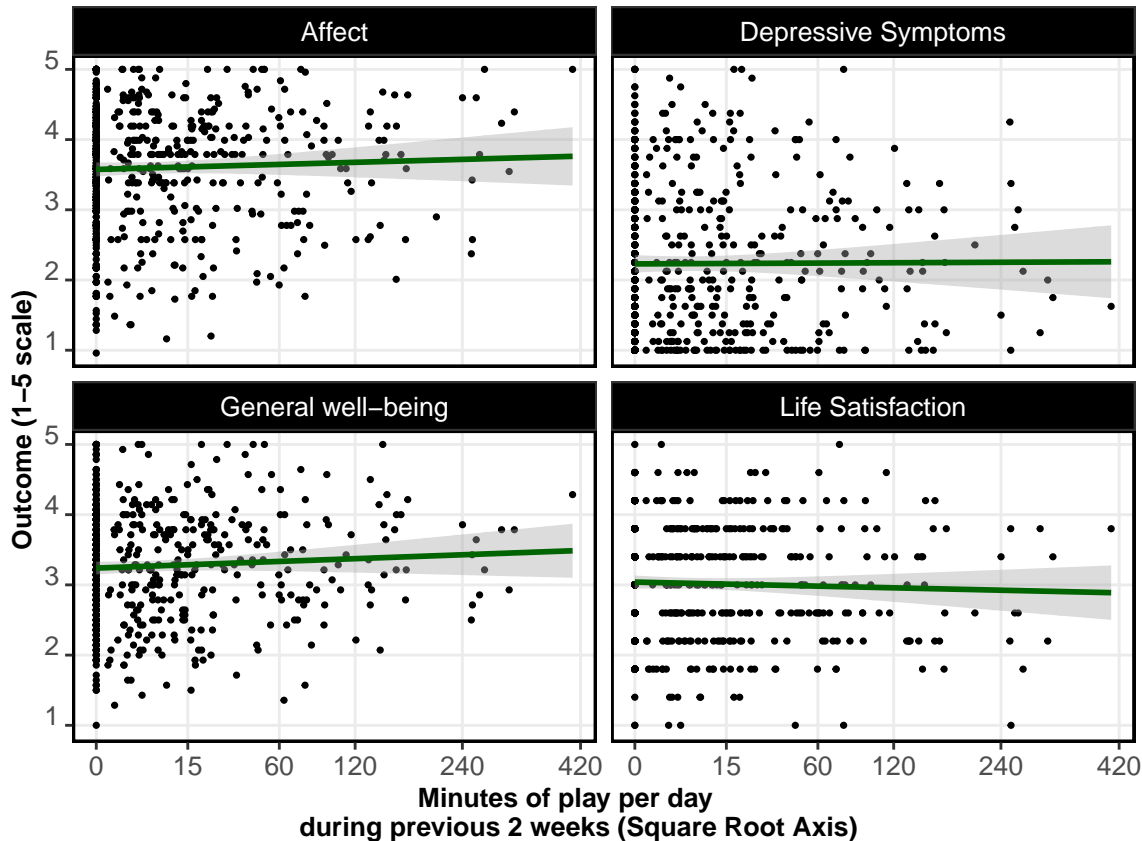
### H2: Exploration of other playtime windows

Next, we conducted exploratory analyses to understand if the relationship between playtime and well-being varies across different playtime periods (Figure 4). Broadly, results align with the results of H1—in all models, 99% CIs overlapped 0, but due to low precision no estimate was fully within the equivalence bounds. We therefore do not find evidence for a meaningful relationship between playtime and well-being at any timescale, but cannot rule out the possibility of one existing.

Estimates are especially uncertain for observation periods of six hours or less, as only 30 participants had played Nintendo games shortly before completing the survey. However, there is a trend towards stronger relationships among more recent observation periods: based on the point estimates, playtime within the previous 1–2 hours is more strongly correlated with well-being than medium- and longer-term time pe-

**Figure 3**

Scatterplots depicting the relationship between video game playtime during the previous 2 weeks (mean minutes of play per day) and four types of well-being.



riods. In each case, playtime shortly before completing a survey was associated with higher affect, life satisfaction, and general mental well-being; and with lower self-reported levels of depressive symptoms.

### Exploratory analysis: Moderation by life fit

Next, we conducted exploratory analyses to investigate what other factors might influence who exhibits positive or negative relations between gaming and well-being. We explored age, gender, and life fit—the perceived harmful or beneficial value of gaming across various life domains outside of play. We expected that people who perceive gaming as supportive in other life domains would show positive relations between gaming and mental health, and those who perceived gaming to be harmful to other life domains would show negative ones.

To test this, we reran the models from H2, adding playtime \* age, playtime \* gender, and playtime \* lifeFit moderation terms. We did not find evidence to

support the presence of moderation; none of the moderation terms were significant ( $0.064 < p < 0.99$ ).

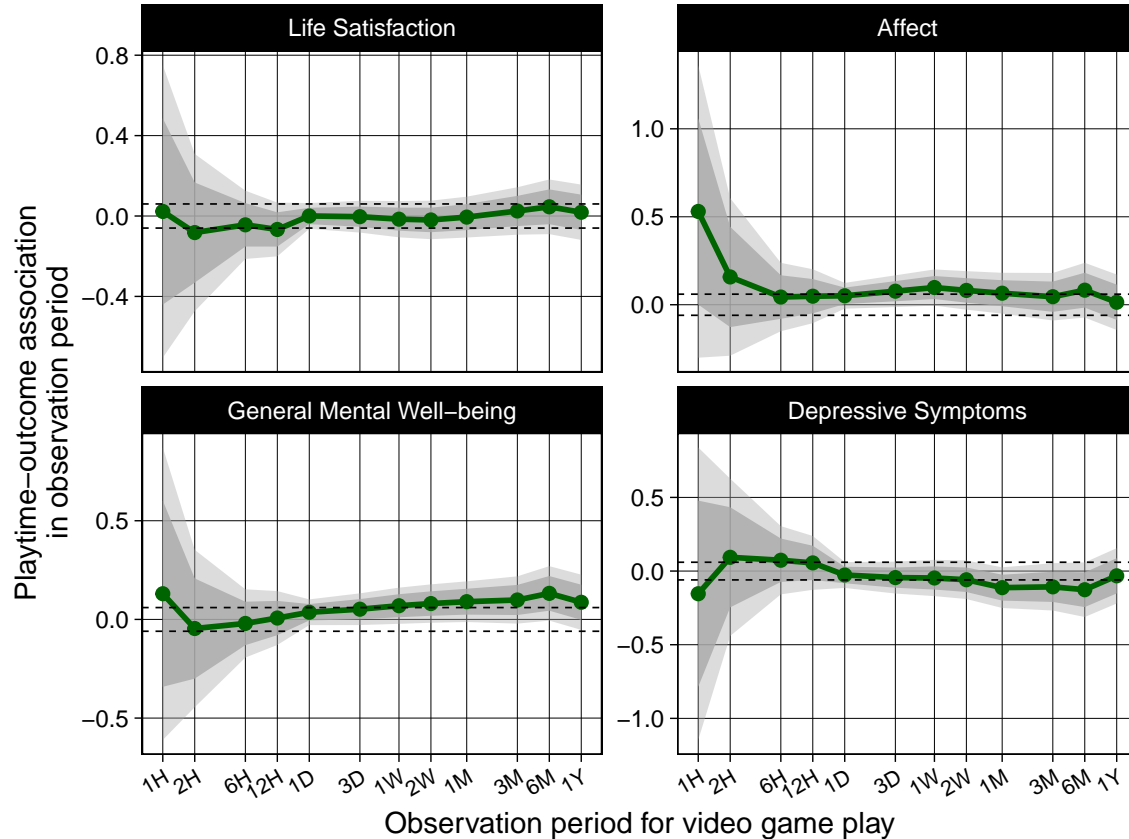
However, we did find evidence of direct positive relationships linking life fit to well-being separate from playtime (Figure 5): Those who believe gaming to be beneficial to their lives tend to be also more likely to report higher levels of well-being, regardless of how much they play. Across 48 models, we observed relationships between well-being and a 1 point change in life fit ranging from 0.153 to 0.321 (median = 0.242; all  $ps < .001$ ).

### Sensitivity Checks

We performed various sensitivity checks to ensure the robustness of our findings, detailed in full in the supplementary materials. First, we explored potential non-linearity in the relationships between playtime and mental health by comparing a generalised additive model of well-being with and without a smooth term for playtime, and comparing AIC between these. Of the 48 possible models (4 well-being vari-

**Figure 4**

Estimates for the relationship between playtime and well-being across various timescales, shown with 90% (dark blue) and 99% (light blue) confidence intervals. Dashed lines represent the positive and negative smallest effect size of interest (SESOI) of .06.



ables \* 12 playtime windows), just one of these (playtime in the previous one year and life satisfaction) showed a difference in AIC of more than two, indicating that nearly all relationships were adequately captured by linear terms. Next, we reran the analyses using session durations as calculated by Nintendo, as opposed to the implied duration based on the start and end timestamps; Nintendo's durations are shorter than the implied duration in approximately 10% of sessions due to internal methodology. Data show a similar pattern: no models showed a significant relationship between playtime and well-being at our specified alpha of .01. Next, we explored alternative models wherein playtime was separated into both a binary variable (1 if the player had any time logged in that period, 0 if not), and a continuous variable (how much a person played). Results were comparable; although three models indicated that among those who played in the previous 1-2 weeks, longer play is associated with higher affect and general mental well-being, the remaining 93 playtime variables were neither significant or nor within both equivalence

bounds.

## Discussion

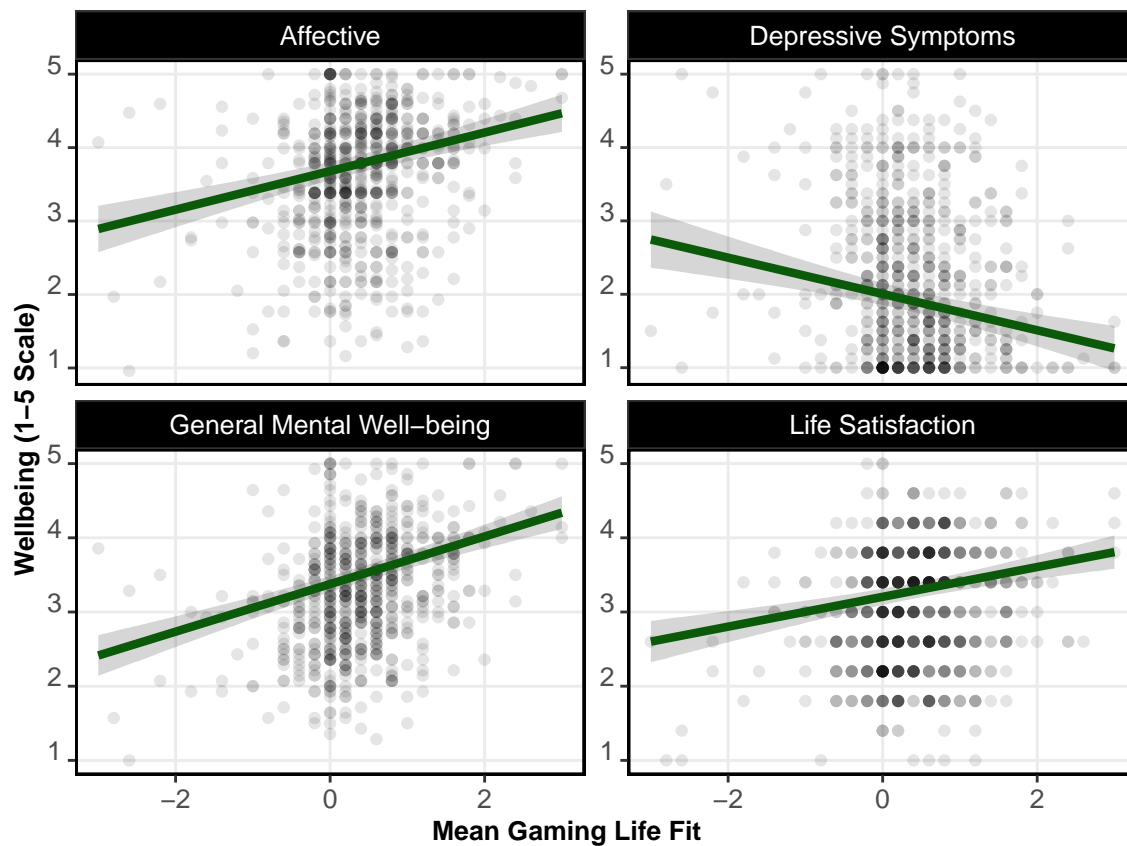
Although we did not intend our study to test the causal question most critical to heated debates surrounding video game engagement and global health, our study is a concrete step in the right direction, having independently recruited a large sample of video game players [as opposed relying on recruitment through games companies themselves; e.g. Johannes et al. (2021)], collected validated measures of well-being, joined these with objective behavioral telemetry and made minimal adjustments for age, gender, employment, and education. Using these methods, we did not uncover robust or consistent relationships between time spent playing and various mental well-being outcomes.

Although not conclusive, our results point toward a pattern whereby platform-wide video game play time does not predict well-being to a meaningful degree. This trend, across a wide



Figure 5

Marginal relationship between gaming life fit (perceived harmful or beneficial effects of games for oneself) and wellbeing



453 range of outcomes, timescales of play, and model specifica-474  
 454 tions adds to a growing body of work that suggests that simple-475  
 455 time spent playing games is unlikely to affect well-being for-476  
 456 the average player. Said differently, the findings we report-477  
 457 place the onus on those who assert that there is a meaningful-478  
 458 relationship between playtime and well-being. It should be a-479  
 459 priority to identify and concretely articulate which confounds-480  
 460 might bias a true effect towards the null associations reported-481  
 461 in this and other research using player telemetry (Ballou et al.,-482  
 462 2024; Johannes et al., 2021; Larrieu et al., 2023; Vuorre et al.,-483  
 463 2022).

464 To further elucidate this point, we conducted brief simu-484  
 465 lation tests to ascertain how strong such confounding might  
 466 need to be (see Supplementary Materials). For example, if-485  
 467 the true standardized effect of playtime on mental health was-486  
 468 a moderate .2 SDs per additional hour of daily playtime, a-487  
 469 confound C would need to be a very strong cause of both-488  
 470 X ( $\beta = .5$ ) and Y ( $\beta = -.5$ ) to bias the true .2 ef-489  
 471 fect to null. While we do not claim this is impossible, we-490  
 472 do believe it unlikely. Approaching the topic along these-491  
 473 lines—identifying confounds, testing the presence or absence-492

of correlations for their sensitivity to potential confounds, and  
 systematically identifying factors that do (not) cause play-  
 time and well-being—can help us achieve more systematic  
 progress (Ballou et al., n.d.). This work can be bolstered by  
 qualitative research aimed at more fully mapping the causal  
 system and by substantive theory development with greater  
 specificity in the aspects of media use expected to produce  
 effects, the hypothesized causal relationships, boundary con-  
 ditions, and so forth (Ballou, 2023; Coenen, 2023; Eronen &  
 Bringmann, 2021; Magnusson et al., 2024).

### Who are “gamers”?

The steady attrition throughout the stages of the research  
 process from screening, to linking, to successful data retrieval  
 highlights the challenges for participant recruitment in video  
 games research. Despite a series of filtering steps wherein a  
 majority of participants were filtered out due to not playing  
 Switch games or being unwilling or unable to link data, our  
 final sample remains only minimally engaged with Nintendo  
 games—playing just 1.4 hours per week on average. As a

493 result, the population here is clearly different than previous 545  
 494 studies recruiting participants with the help of games compa-546  
 495 nies (Johannes et al., 2021; Larrieu et al., 2023) or through 547  
 496 social media forums for highly-engaged players (Ballou et al., 548  
 497 2024). 549

498 As argued above, we believe this is a valuable group in its 550  
 499 own right—those who rarely play video games may be partic-551  
 500 ularly susceptible to their positive or negative effects on the  
 501 occasions they do play. While the current study is unlikely to 552  
 502 generalize to so-called “hardcore” players who play several  
 503 hours per day or more and therefore may experience more 553  
 504 accumulative effects, our findings align with previous stud-554  
 505 ies of more highly engaged players and add a new subgroup 555  
 506 of players to the body of work showing the absence of any 556  
 507 meaningful relationship (Ballou et al., 2024; Larrieu et al., 557  
 508 2023; Vuorre et al., 2022). 558

509 As the field progresses, however, differences in the level 559  
 510 of engagement pose major challenges for study sampling and 560  
 511 generalisability. Calls for representative samples need to 561  
 512 specify the population of reference: should this be the general 562  
 513 population (of whom many do not play games), people who 563  
 514 play any games at all (of whom many do not play the games 564  
 515 for which researchers have data access), people who play the 565  
 516 particular game or platform of interest (of whom many may 566  
 517 be only minimally engaged), or something else entirely? In 567  
 518 the field’s quest for more generalizable results, this will be a 568  
 519 critical issue. 569

## 520 Timescales 571

521 While we are quick to caution that this is a preliminary 573  
 522 finding that should not be relied upon without further vali- 574  
 523 dation, our data provide some initial evidence that effects of 575  
 524 raw playtime might materialize and fade within a few hours— 576  
 525 point estimates indicated that playtime was more strongly 577  
 526 linked with greater well-being in the 1-2 hours prior to survey 578  
 527 completion. This finding is compatible with various causal 579  
 528 explanations: for example, players who recently played are 580  
 529 more likely to be in a period of leisure time, which would be ex- 581  
 530 pected to generate more positive feelings than in peers doing  
 531 obligatory activities such as work. However, if researchers do 582  
 532 expect to see positive effects of gaming, our data suggest that  
 533 they may need to search for highly proximal effects directly 583  
 534 during and following a play session (e.g., Vuorre et al., 2024). 584

535 Should this finding be upheld, it would go a long way 585  
 536 towards explaining previous null findings from studies that 586  
 537 related well-being to playtime over timescales such as two 587  
 538 weeks (Vuorre et al., 2022), one month (Sibilla et al., 2021), 588  
 539 six months (Weinstein et al., 2017), and one year (Kowert et 589  
 540 al., 2015). For most players, it may be the case that gaming is 590  
 541 a recovery activity that helps to manage day-to-day stresses 591  
 542 and mood fluctuations, without necessarily having substan- 592  
 543 tial long-term impacts. The majority of players have several 593  
 544 options for activities in their environment that would have 594

comparable effects on their well-being. Such activities are  
 thus “exchangeable”, serving the same short-term goals with-  
 out consequences for long-term adjustment. Studying rela-  
 tionships over the course of hours has to date been possible  
 largely only in laboratory settings—rarely have researchers  
 had access to session-level data of naturalistic behavior that  
 they could link to momentary well-being.

## Life fit

This study demonstrates the potential usefulness of life fit  
 as a theoretical construct (Ballou & Deterding, 2023). Given  
 the accumulating evidence that playtime and well-being are  
 not meaningfully related at the population level alongside in-  
 controvertible evidence that some players benefit and some  
 are harmed (Ballou et al., n.d., n.d.), the task for the field can  
 be framed as a search for the most important moderators. Life  
 fit—a player’s self-assessment about the contribution of gam-  
 ing to different aspects of their lives—stands as an effective  
 starting point, letting researchers trust the lived experiences  
 of players to guide them towards patterns of problematic or  
 particularly beneficial play.

Using this measure, we found no evidence that life fit mod-  
 erated the relationship between playtime and well-being, but  
 we did find a direct correlation between the two. Notably, this  
 relationship was an order of magnitude stronger than any es-  
 timates for playtime itself. Among several other possible ex-  
 planations, this would fit a pattern of biased self-assessment:  
 it is possible that players who are generally feeling poorly  
 are more likely to appraise their gaming as harmful to their  
 mental health, regardless of whether that mechanism actually  
 takes place. This would align with some previous findings  
 that more depressed people tend to overestimate their smart-  
 phone use due to a negative or guilt-laden appraisal process  
 distinct from the media use itself (Sewall & Parry, 2021).

We caution that the measure applied here has not been val-  
 idated, and is better viewed as a formative indicator than as a  
 true latent variable. More work will be needed to understand  
 the validity of this construct.

## Holistic digital trace data

This paper demonstrates both the value and difficulty of  
 collecting holistic digital trace data: by capturing data across  
 an entire platform, rather than just one game, we can poten-  
 tially account for a person’s complete engagement with games  
 without self-report biases—but only if we sample players for  
 whom that platform constitutes the majority of their gam-  
 ing. Our screening data indicates that participants play games  
 on average across 2.8 platforms, for example playing games  
 across Nintendo, Steam, and iOS. To fully capture players’  
 entire gaming diets, researchers will need to either subsam-  
 ple participants who use only one platform or develop distinct  
 methods of collecting digital trace data for several platforms.

In this same vein, although collaborations between academics and digital media platforms are becoming increasingly more common (Larrieu et al., 2023; e.g., Nyhan et al., 2023), these remain difficult to source and stubbornly inequitably distributed across the research ecosystem. Researchers are actively exploring other ways to source digital trace data, including through scraping methods (Ballou et al., 2024), APIs (e.g., <https://gameplay.science>), and subject data access requests/data download packages (Breuer et al., 2022) but more needs to be done. Relationships between game firms and independent research teams are not scalable and the providence of data collected by scraping and related tools is difficult if not impossible to verify. Democratizing researcher access in a way that protects participants' autonomy and right to privacy will require the enactment of multisector-spanning initiatives like the UK's Video Game Research Framework (Department for Media & Sport, 2023) that clearly prescribe the responsibilities for those enabling, enacting, and benefiting from the scientific study of video game play. The time for this is well past due.

### Limitations

There are three limitations and constraints on generalisability the merit mention. First, we could not analyse telemetry generated when players engaged with third-party titles (i.e. games not published by Nintendo or its closely associated companies). As a result, our findings only hold for similar Nintendo games (75% of which were rated for everyone or everyone over 10), and it is not possible to rule out the possibility that third-party games with different content or themes might show a different pattern of effects. Likewise, because we collaborated with Nintendo of America, our sample consisted only of adults living in the United States, a group that we found to be largely casually involved with first-party Nintendo games. Games are both global and played by those of all ages, so it is not clear the degree to which our findings do or do not apply to younger players, those who play other games, or those who approach games from different cultural and linguistic backgrounds. Finally, while we have longitudinal telemetry data, our self-report survey was cross-sectional and this study is one step towards designs using daily diary and experience sampling methods (Aalbers et al., 2021; e.g., Siebers et al., 2021).

### Conclusion

The idea that time spent playing is the key ingredient in how games impact well-being will be with us for some time. Although our study was not designed to test a causal link, it challenges the notion that simply playing more affects well-being, for better or for worse. The correlations we observed were mostly too small to practically matter. Moreover, we show that profound confounding would be required to suppress a true causal effect to account for the null associations

we report. This is improbable but not impossible, and we believe our results lend weight to calls for scholars and health practitioners to embrace the gradual shift towards focusing on the quality, rather than quantity of video game play as the key factor for player health. If this can be done while simultaneously improving data quality and access, a coherent and evidence-based method for studying the complex relationships linking video game play and well-being will be possible.

### Contributions

Author roles were classified using the Contributor Role Taxonomy (CRediT; <https://credit.niso.org/>) as follows: Nick Ballou: conceptualization, data curation, methodology, formal analysis, writing; Matti Vuorre: methodology, funding acquisition, formal analysis, editing; Thomas Hakman: data curation, validation, editing; Kristoffer Magnusson: methodology, editing; Andrew K Przybylski: conceptualization, funding acquisition, project administration, editing

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